Mapping the Spatial Scales of Australian Extreme Precipitation Using Daily Rain Gauges

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Abstract

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The impacts of extreme precipitation events (EPEs) on society are strongly influenced by their spatial footprint, yet the spatial scales of such events remain underexplored. Here we present the first continent-wide analysis of the spatial scales of daily EPEs in Australia. We estimate the characteristic spatial scale of EPEs seasonally across the Australian continent using daily station observations and semivariograms. A semivariogram is a spatial statistical function that measures how spatial autocorrelation in precipitation decays with distance. Consistent with global analyses of satellite data, EPEs generally have larger spatial scales at higher latitudes. However, our analysis reveals complex seasonal and geographical dependencies that highlight the role of topography and meteorological regimes. We also analyse EPE spatial scales under different phases of the El Niño-Southern Oscillation (ENSO). In SON and DJF, southeastern Australia exhibits larger spatial scales during La Niña, although no uniform pattern is observed across the continent. Long-term changes were analysed using 2,070 stations with continuous operation between 1960 and 2023. Southwestern Australia shows a notable reduction in median EPE length scale in most seasons, while eastern regions exhibit a decrease in MAM and an increase in SON. Together, these findings provide a new climatological reference for the spatial scale of EPEs in Australia. These results also highlight the need to better understand the physical factors controlling the spatial scale of precipitation extremes in current and future climates.

1 Introduction

The impacts of extreme precipitation events (EPEs) are strongly modulated by their spatial footprint (e.g. Schaller et al., 2016; Touma et al., 2018; Bevacqua et al., 2021, and citations therein). For example, widespread EPEs can lead to spatially extensive floods affecting multiple regions simultaneously, amplifying societal impacts across broader geographic areas (Jongman et al., 2014; Berghuijs et al., 2019). In contrast, localised extremes may cause intense but spatially confined impacts, such as flash flooding, which often overwhelm local infrastructure and are harder to predict due to short lead times and high spatial variability (e.g. Schumacher, 2017). Although many studies have focused on the frequency and intensity of EPEs, relatively few have explored their spatial extent (Hoegh-Guldberg et al., 2018; Bevacqua et al., 2021).

The characteristics of EPEs are expected to change in a warming climate (O'Gorman, 2015). As the atmosphere warms, it becomes moister, favouring more intense precipitation, while shifts in large-scale circulation patterns may alter both precipitation intensity and spatial patterns (O'Gorman & Schneider, 2009; Bevacqua et al., 2020). Several studies have examined how the spatial extent of EPEs responds to warming, but no clear consensus has emerged from modelling and observational approaches. In modelling studies, Chang et al. (2016) reported a reduction in storm size under a warmer climate, whereas Guinard et al. (2015) and Dwyer and O'Gorman (2017) found that the precipitation structures increase in size with warming. Observational analyses of the present-day climate have also yielded different findings for the relationship between rainfall spatial extent and atmospheric conditions: Wasko et al. (2016) observed a decrease in storm size with increasing local temperatures using hourly observations in Australia, while Lochbihler et al. (2017) identified a positive relationship between rain cell size and dew point temperature in the Netherlands. In contrast, a global trend analysis by Tan et al. (2021) found that the spatial extent of EPEs increased from 1983 to 2018 in non-monsoon regions of the Northern Hemisphere.

Beyond the divergent findings across existing studies, the spatial scales of extreme precipitation remain underexplored. Australia is one of the many regions that still lack a baseline characterisation of the spatial extent of EPEs in the current climate. While Australian rainfall is "more variable than could be expected from similar climates else-

where in the world" (Nicholls et al., 1997), many existing studies of precipitation extremes focus on specific subregions without explicitly addressing spatial scales (e.g., Warren et al., 2021; White et al., 2022). Some limited analysis has been done at the hourly level by Wasko et al. (2016). They estimated storm sizes using hourly station data from a small number of stations (93 for one-hour events and 78 for three-hour events) in Australia, thereby restricting their spatial representativeness. On the daily scale, Saunders et al. (2021) used annual maxima to identify regions that are likely to be similarly affected by EPEs for Australia. However, annual maxima at different locations do not necessarily occur on the same day, which limits their utility for analysing the spatial extent of extremes.

To address this existing gap, this study provides a climatological assessment of EPE spatial extent across the Australian continent using daily rain gauge data. A well-defined climatology enables an assessment of the mean structure and variability of EPEs. This provides a foundation for investigating the physical mechanisms that govern their spatial extent. Long-term gauge measurements also allow us to evaluate how EPE spatial scales respond to internal climate variability and provide estimates of long-term trends.

Although a global EPE spatial scale analysis was recently developed by Tan et al. (2021) using satellite observations, our analysis relies on in situ rain gauge data. While gauge-based estimates are subject to measurement errors and are constrained by limited spatial coverage, they ensure greater accuracy in capturing precipitation extremes and provide the advantage of long observational records. Satellite-based products often exhibit precipitation estimation errors when compared to ground measurements. For example, Tansey et al. (2022) found that satellite products tend to overestimate seasonal total liquid precipitation; Montoya Duque et al. (2023) reported an overestimation of precipitation intensity in the Southern Ocean; and Ponukumati et al. (2023) showed that satellite-based products can overestimate extreme rainfall intensity relative to daily automatic weather station measurements in India. These errors limit the reliability of satellite products for analysing precipitation extremes. Compared to rain gauges, satellite products are also suboptimal for long-term analyses of extreme events. Most satellite-based precipitation missions were launched after the 1980s (e.g., Simpson et al., 1996; Ashouri et al., 2015), and typically provide only one or two observations per day, while many rain gauges in Australia were deployed in the early 1900s (see Figure S1).

To estimate spatial scales, we adopt the semivariogram approach introduced by Touma et al. (2018). Originating in spatial statistics, a semivariogram describes how the spatial dependence in a variable, such as precipitation, weakens as the distance between two locations increases. It provides an estimate of the distance over which the variable remains correlated (Cressie, 2015). This estimate serves as a quantitative indicator of the spatial scales of EPEs. This semivariogram approach was first used to characterise the spatial scales of EPEs across the United States using the Global Historical Climatology Network-Daily (GHCN-D) station dataset (Touma et al., 2018). It was further extended by Touma et al. (2019) to assess the spatial extent of tropical cyclones, and more recently adopted for the global satellite-based EPE spatial scale study by Tan et al. (2021).

Building on the existing semivariogram-based work, we use all stations available from the Bureau of Meteorology's (BoM) daily rain gauges (9861 in total) to derive a long-term climatology of the EPE spatial scales for the Australian continent. Since ENSO is the main driver of interannual rainfall variability in northern and eastern Australia (e.g. McBride & Nicholls, 1983; Risbey et al., 2009; King et al., 2020; Gillett et al., 2023; Tozer et al., 2023; McGregor et al., 2024; Huang et al., 2024; He et al., 2025), we further examine how the EPE spatial scale differs between El Niño and La Niña periods. Finally, we leverage the long-term record of daily rain gauge observations to assess changes in EPE spatial scales over time. Our findings show that the seasonal patterns, ENSO relationships, and long-term trends in EPE spatial scales vary across the continent but differ from those of mean precipitation and those of the intensity of precipitation extremes.

In particular, increases in spatial scale do not consistently occur during "wetter" periods or in "wetter" regions. This suggests that different physical processes govern the intensity and spatial extent of precipitation extremes, and that these characteristics may respond differently to climate change.

The remainder of this paper is structured as follows. Section 2 describes the rain gauge dataset, the semivariogram approach used to estimate the EPE spatial scales, ENSO classification, and trend analysis methodology. Section 3 presents our results on the spatial scales of EPEs, including their regional and seasonal variability, relationship with ENSO, and long-term trends. Sections 4 and 5 provide discussions and conclusions, respectively.

2 Data and Methods

2.1 Data for analysing seasonal EPEs

We used daily rain gauge observations obtained from the weather station directory operated by the BoM for the period between 1940 and 2024. The BoM daily rainfall observations are made at 0900 local time, recording the preceding 24-hour total precipitation. The earliest rain gauge record in the BoM archive started in May 1826. The number of active BoM stations in Australia over time is shown in Figure S1.

To estimate the climatological length scales of EPEs, we need stations that have continuous data availability, a sufficiently long period of record, and satisfactory data quality and station density. To balance these factors, we chose the period from 1940 to 2024 for analysis. Only stations with an operation period of more than 20 years were included. A total of 9,861 stations met this criterion between 1940 and 2024. The locations of these stations are shown in Figure 1a.

The stations are grouped using the fifteen natural resource management (NRM) sub-clusters, which are administrative regions used for the Australian Government's environment and sustainable agriculture programs (CSIRO, 2015). These regions provide a standardized geographic framework widely used in environmental assessments across Australia. Considering the climate characteristics and spatial station distribution in the rangelands (central Australia), we divide the rangeland region into western (RW) and eastern (RE) clusters along the 130°E longitude, replacing the original north and south rangeland clusters from the NRM classification (see CSIRO, 2015). The acronyms for the NRM sub-clusters are given in Table 1.

To identify daily extremes, we define each station's seasonal 90th percentile of rainy days (precipitation > 1 mm day⁻¹) as its extreme precipitation threshold (hereafter "P90"). Conditioning on rainy days avoids including weak rainfall in dry regions and prevents clustering events in a single season. An EPE is recorded when daily precipitation meets or exceeds the station's seasonal P90. We used the seasonal P90 instead of the monthly P90 as in Touma et al. (2018) and Tan et al. (2021) because this better reflects seasonal climatology and avoids including moderately heavy events from drier months in the analysis.

The P90 values across each season and each station are shown in Figure 2. Hereafter, to avoid overplotting in regions with high station density, station data are averaged into 40-km grid boxes. The box size was chosen for visualization purposes, providing a balanced representation of both dense and sparse regions. The seasonal variation in P90 magnitudes reflects well-known rainfall patterns across Australia, with the highest values occurring in DJF and much lower values in JJA (e.g. Drosdowsky, 1993). This seasonality is particularly pronounced in Northern Australia, where the monsoon brings most rainfall between October and April. During the monsoon inactive phase (May–September), most JJA days in Northern Australia are dry, and rainfall amounts on wet days are low,

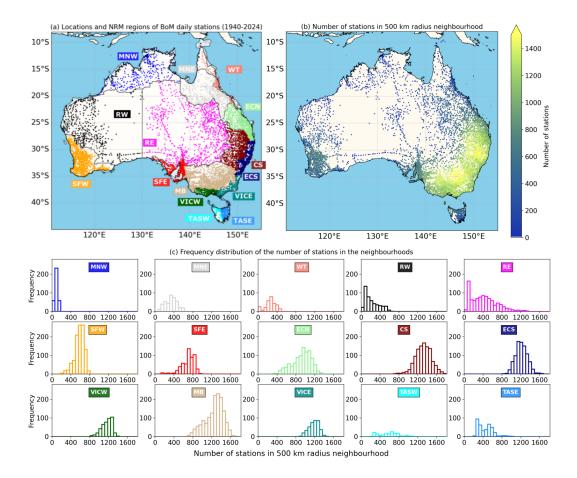


Figure 1. (a) Locations of BoM daily stations with more than 20 years of continuous records between 1940 and 2024 coloured by their NRM regions. (b) The average number of stations used in each 500-km radius neighbourhood for P90 semivariogram analysis. (c) Frequency distribution of the number of stations in the 500-km radius neighbourhoods in each NRM region, as shown in Panel (a).

resulting in a lower P90 magnitude. These variations are not evident in the global study by Tan et al. (2021), possibly because their satellite study examined spatially averaged precipitation over $0.25^{\circ} \times 0.25^{\circ}$ regions rather than gauge measurements.

2.2 Estimating the spatial scales of EPEs

2.2.1 Data preparation for spatial scale estimation

To estimate the spatial scale of EPEs, the daily precipitation data were first converted into a binary EPE dataset. At each station, observations were assigned a value of 1 if their daily precipitation is greater than or equal to the seasonal P90 threshold, and 0 otherwise. We did not retain precipitation amounts because the magnitude of extremes is relative to geographical context and can vary with measurement instrumentation, especially across a continent as large as Australia. By using binary data, the impacts of extreme precipitation events can be compared across regions, whereas raw precipitation amounts vary greatly across Australia (see Figure 2).

To identify stations impacted by the same EPE, we define a 500-km radius neighbourhood around each station (see black and red dots in Figure 3a, c). On days when

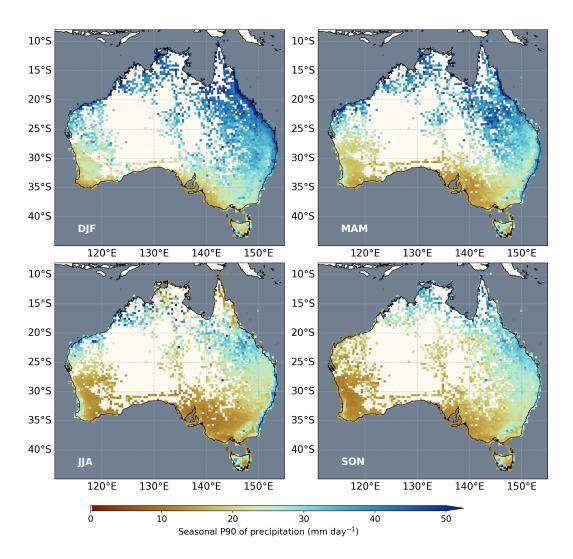


Figure 2. Seasonal 90th percentile (P90) daily precipitation accumulation for DJF (top left), MAM (top right), JJA (bottom left), and SON (bottom right)

Table 1. Acronyms for the NRM sub-clusters shown in Figure 1. Note the Rangeland has been split into Rangeland West and Rangeland East due to the natural separation of station locations.

Acronym	NRM Sub-Clusters
MNW	Monsoonal North West
MNE	Monsoonal North East
WT	Wet Tropics
RW	Rangeland West
RE	Rangeland East
SFW	Southern Flatlands West
SFE	Southern Flatlands East
ECN	East Coast North
CS	Central Slopes
ECS	East Coast South
VICW	Victoria West
MB	Murray Basin
VICE	Victoria East
TASW	Tasmania West
TASE	Tasmania East

the neighbourhood central station exceeds its seasonal P90 threshold, all stations within this neighbourhood are paired and classified into the following categories:

1. 1-1 pair: two stations both detected an EPE on the same day.

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- 2. 1-0 pair: only one of the two stations detected an EPE on that particular day.
- 3. 0-0 pair: the EPE was observed at neither station on that day.

The 0-0 pairs are discarded as they are not needed in the subsequent analysis, similar to Touma et al. (2018). Figures 3a and 3c demonstrate these pair categories: any two red stations are classified as a 1-1 pair; a red and a black station as a 1-0 pair; and two black stations as a Pair 0-0.

This approach of pair selection does not account for EPEs that occur beyond the neighbourhood boundary. To address this, we allowed for relaxed neighbourhood pairing, meaning that one of the stations in a pair could lie outside the 500-km neighbourhood. This is illustrated in Figure 3a and 3c (pink and grey dots). Only pairs where at least one station lies within the 500-km neighbourhood (i.e., red-pink, red-gray, and black-pink pairs) were retained. Pairs entirely outside the 500-km radius were excluded.

The choice of neighbourhood size is a key consideration, as previous studies have shown it can influence the absolute magnitude of estimated spatial scales, although the relative regional and seasonal variations remain largely unchanged (Touma et al., 2018). On this basis, we conservatively adopt a 500-km radius, consistent with physical reasoning that this scale is sufficient to capture the spatial extent of many mesoscale systems (e.g., Khouakhi et al., 2017; Tan et al., 2021). The number of stations used for EPE spa-

tial scale analysis within the 500-km neighbourhoods for each station and NRM region is summarised in Figure 1b-c.

2.2.2 Semivariogram estimation using the binary EPE dataset

To quantify the EPE spatial scales, a semivariogram is used to characterise the degree of spatial autocorrelation of a given spatial field Z(x) at location x by describing how the similarity between observations decreases as the distance h between them increases. It is defined as the expected squared difference between values separated by distance h:

$$\gamma(h) = \frac{1}{2} \mathbb{E} \left[Z(x) - Z(x+h) \right]^2, \tag{1}$$

where h is the distance between the two locations (Webster & Oliver, 2007). This form of the semivariogram assumes isotropy, which means that the spatial relationships depend only on the distance between locations, not on their spatial orientation relative to one another.

As the EPE dataset is now binary, Z(x) only takes values of 0 or 1. This means the pointwise estimate of $\gamma(h)$ for any two stations x_i and x_j is 0 for a 1-1 pair, and 0.5 for a 1-0 pair. To empirically estimate $\gamma(h)$, we require an approach suited to binary data, irregularly spaced stations, and changing network coverage.

To address this, the observations can be pooled to estimate the semivariogram empirically. We group station pairs (1–1 and 1–0) based on their distance apart, using 25-km bins within a 500-km neighbourhood, following Touma et al. (2018) and Tan et al. (2021). The grouping intervals are therefore (0,25] km, (25, 50] km, ... (475, 500] km. For any interval $(h - \delta, h]$, where $\delta = 25$ km is the bin size, this can be written as:

$$\mathcal{N}_h = \{ (x_i, x_j) : h - \delta < d(x_i, x_j) \le h; i, j = 1, \dots S \},$$
(2)

where S is the total number of stations in the relaxed neighbourhood, $d(x_i, x_j)$ is the Euclidean distance between stations x_i and x_j . This produces a series of sets containing grouped station pairs $\mathcal{N}_{25}, \mathcal{N}_{50}, \ldots \mathcal{N}_{500}$.

By changing the expectation in Equation 1 to a sum, the empirical estimate of the semivariogram for a given binned interval is

$$\hat{\gamma}\left(h - \frac{\delta}{2}\right) = \frac{1}{2|\mathcal{N}_h|} \sum_{(x_i, x_j) \in \mathcal{N}_h} \left[Z(x_i) - Z(x_j)\right]^2,\tag{3}$$

where $h - \frac{\delta}{2}$ is used because estimates are taken at the bin centre. The notation $|\mathcal{N}_h|$ means the total number of distinct station pairs in the set \mathcal{N}_h . As the pairs fall into one of two distinct categories (1-0 and 1-1 pairs), this can also be written as $|\mathcal{N}_h| = |\mathcal{N}_{h,(1,0)}| + |\mathcal{N}_{h,(1,1)}|$.

As Z(x) is a binary field, the above equation further simplifies to

$$\hat{\gamma}\left(h - \frac{\delta}{2}\right) = \frac{1}{2}\left(\frac{|\mathcal{N}_{h,(1,0)}|}{|\mathcal{N}_h|}\right) = \frac{1}{2}\left(\frac{|\mathcal{N}_{h,(1,0)}|}{|\mathcal{N}_{h,(1,0)}| + |\mathcal{N}_{h,(1,1)}|}\right). \tag{4}$$

Intuitively, this estimator represents the strength of the spatial relationship, which can be estimated as the ratio of 1-0 pairs to the total number of pairs in that binned interval. Visually, this is shown in Figure 3b, d, e, where the empirical estimates (black

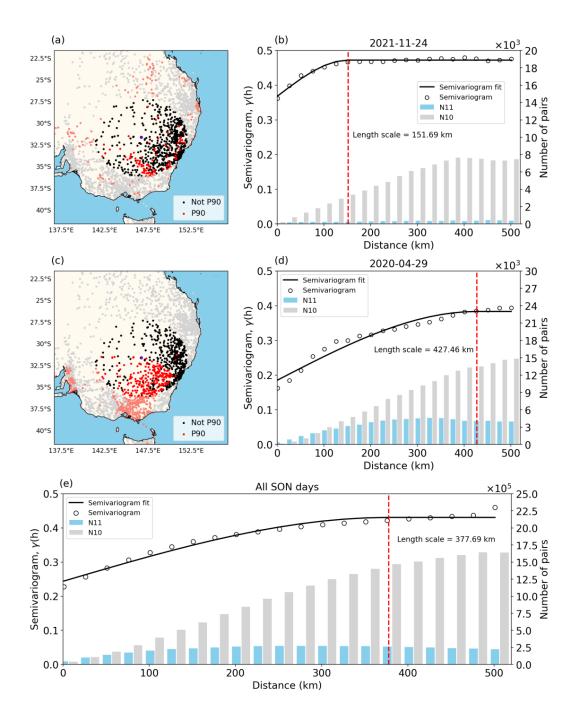


Figure 3. Illustration of the relaxed moving neighbourhood method for a station in New South Wales, Australia. Two examples are given for the EPE on 24th November 2021 (a-b), and a larger EPE on 29th April 2020 (c-d). Panels (a) and (c) are maps illustrating stations within the 500-km radius neighbourhood in red (P90 detected) and black (P90 not detected). Stations outside the neighbourhood are shown in pink (P90 detected and outside) and grey (P90 not detected and outside). Panels (b) and (d) show the semivariogram estimation for the two days. The blue bars represent the number of 1-1 pairs, and the gray bars are for 1-0 pairs. The red dashed line shows the derived length scale. Panel (e) illustrates the climatological length scale for the station in SON. The blue and gray bars are the sum of all 1-1 pairs and 1-0 pairs over the 67 extreme days in SON for that station, respectively.

dots) correspond to half the ratio of the gray bar height to the combined height of the gray and blue bars.

Figures 3a-d illustrate semivariogram estimation for two individual days. To ensure more robust estimation, here we focus on estimation of a climatological semivariogram $\hat{\gamma}_{season}(h)$ based on all EPEs within a given season. To obtain climatological length scales, we grouped all 1-1 pairs and 1-0 pairs within each station's 500-km neighbourhood based on the bin interval $(h-\delta,h]$, pooling across all EPE days (d) within a given season from 1940 to 2024:

$$\hat{\gamma}_{season}\left(h - \frac{\delta}{2}\right) = \frac{1}{2} \left(\sum_{d \in \mathcal{D}} \frac{\left| \mathcal{N}_{h,(1,0)}^{(d)} \right|}{\left| \mathcal{N}_{h}^{(d)} \right|} \right) = \frac{1}{2} \left(\sum_{d \in \mathcal{D}} \frac{\left| \mathcal{N}_{h,(1,0)}^{(d)} \right|}{\left| \mathcal{N}_{h,(1,0)}^{(d)} \right| + \left| \mathcal{N}_{h,(1,1)}^{(d)} \right|} \right), \tag{5}$$

where \mathcal{D} is the set of EPE days for the season of interest. This process is illustrated in Figure 3e. The same procedure was applied to estimate spatial scales of EPEs associated with ENSO in each season (see Section 2.3).

For this estimator to be robust, we required each station to have at least 20 neighbours within its 500-km neighbourhood, excluding 14 stations that did not meet this criterion. We further retained only EPE days where at least 10% of neighbourhood stations exceeded P90, thereby filtering out events detected by only one or two stations. A sensitivity test without this filter yielded similar overall results, but isolated-station events lacked sufficient spatial coherence for reliable scale estimation.

The semivariogram in this form (Equations 4 and 5) does not reveal the spatial scale of EPEs. To estimate the spatial scale of EPEs, we next fit a parametric model to the semivariogram estimate $\hat{\gamma}(h)$ as described below.

2.2.3 Estimating semivariogram parameters

Various theoretical semivariogram models can be used to model spatial autocorrelation (Webster & Oliver, 2007; Cressie, 2015). For example, Touma et al. (2018) and Tan et al. (2021) used the exponential model, while Touma et al. (2019) adopted the spherical model. Using synthetic data, we compared exponential and spherical fits, finding heuristically that the spherical model was less sensitive to changes in station coverage and exhibited less bias than the exponential model. Based on this, we adopt the spherical model for semivariogram fitting and leave it to future work to investigate parametric model selection further. The spherical semivariogram is given by

$$\gamma(h) = \begin{cases} 0, & h = 0\\ c + b\frac{3h}{2\alpha} - \frac{1}{2}(\frac{h}{\alpha})^3, & 0 < h \le \alpha\\ c + b, & h \ge \alpha \end{cases}$$
 (6)

where c, b, and α are the nugget, the partial sill, and the length scale of the semi-variogram, respectively (Cressie, 2015). The nugget c accounts for variability at very short distances. The nugget and partial sill determine the sill (b+c), which is the $\gamma(h)$ value approached by the spherical semivariogram at large separation distances. The length scale α is the separation distance at which the semivariogram levels off, indicating the distance beyond which spatial proximity no longer governs the semivariogram (red dashed lines in Figure 3b, d, e). Following Touma et al. (2018), we use the length scale α of $\hat{\gamma}_{\text{season}}$ to represent the climatological spatial scale of EPEs. The fitted semivariogram was calculated using the GSTools package with the Python programming language (Müller et al., 2022).

2.3 ENSO events classification and bootstrapping

To identify ENSO EPEs, we use the monthly Southern Oscillation Index (SOI). The SOI is based on the standardized anomaly of the mean sea-level pressure differences between Tahiti and Darwin, with records available from January 1876 (Trenberth, 1984). Sustained negative SOI values below -7 are typically associated with El Niño conditions, while sustained positive values above 7 indicate La Niña conditions.

To ensure adequate and comparable samples for semivariogram estimation, stations must record at least 20 EPEs in both ENSO phases within a season. Thus, only stations with \geq 20 El Niño and \geq 20 La Niña events per season are retained (e.g., a station with 5 El Niño and 30 La Niña events in SON is excluded). To have sufficient stations meet this criterion, the SOI threshold was relaxed from ± 7 to ± 5 . The EPEs are then classified by their monthly SOI with SOI > 5 as La Niña and SOI < -5 as El Niño. After relaxing the SOI thresholds, 4041 out of the total 9861 stations meet the minimum event number requirement in at least one season. The semivariogram estimation was done using all 9861 available neighbour stations with 500-km neighbourhoods centred at these 4041 stations.

To quantify uncertainty in the estimated length scales, we applied bootstrapping separately for each ENSO phase and season at each station. While various bootstrapping methods are available (see e.g. Paciorek et al., 2018, and citations therein), we used the basic bootstrap for simplicity. To equalize sample sizes between phases, the number of events used in bootstrapping was set to the smaller of the two phases within a given season. For example, if a station had 60 La Niña and 90 El Niño events in DJF, we first ranked the 90 El Niño events by the absolute SOI and then subsampled 60 events evenly across the ranked distribution, thereby preserving coverage across the SOI range. For both phases, the 60 events were then resampled with replacement for bootstrapping. The with-replacement sampling was repeated 1000 times. For each bootstrap sample, the semi-variogram approach was applied to estimate the climatological length scale, producing distributions of length scales for El Niño and La Niña events, as well as their differences (La Niña minus El Niño).

To compare EPE length scales between ENSO phases, we used the differences in the median values of their respective bootstrap samples. Statistical significance was assessed at the 90% confidence level based on bootstrap distributions. We consider differences in the median statistically significant if at least 90% of the bootstrap samples are on the same side of zero as the differences in the median. More specifically, significance is assigned when:

- the differences in the median and the 10th percentile of the 1000 sample differences are both positive or,
- the differences in the median and the 90th percentile of the 1000 sample differences are both negative.

2.4 Trend analysis using long-term rain gauge records

The long time coverage of the data used in this study naturally raises the question of how the length scales of EPEs change over time. Unfortunately, the number of active BoM stations varies over time, with many stations being decommissioned and new stations coming online (see Figure S1). Only 1705 stations in Australia have been continuously operating from 1940 to 2024, and only a few of these stations have enough neighbour stations for semivariogram analysis. Therefore, for the trend analysis, we only present the climatological length scales of EPEs for the 500-km neighbourhood centred at the 2070 stations that have been operating from 1st January 1960 to 31st December 2023, which we refer to as "long-term stations" (locations see Figure S2). This is to balance the need between the spatial coverage and the length of the record. The climatological

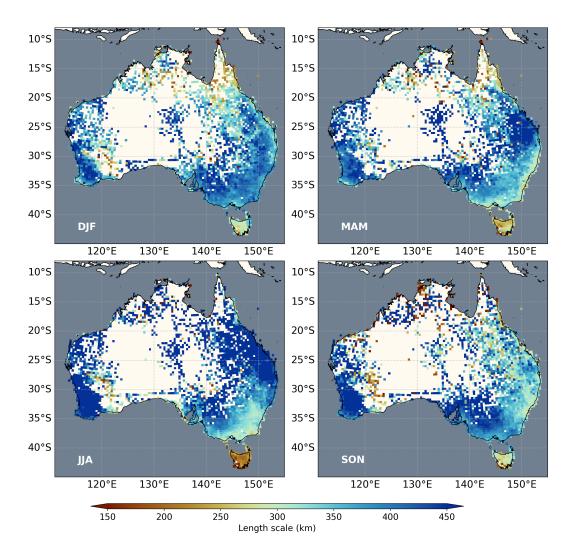


Figure 4. Seasonality of the climatological length scale of EPEs based on 40-km gridded station averages.

length scale results for the long-term stations generally agree with those for the 9861 stations presented below (see Figure S3).

For the trend analysis, we first divided the period between 1960 and 2023 into two halves: 1960-1991 and 1992-2023. Then, we estimated the climatological length scales separately for the two periods for each long-term station. Finally, we compare the climatological length scales for each season between the two periods to assess how the EPE length scales change over time.

3 Results

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3.1 Seasonal length scale of EPEs

Figure 4 shows the seasonal climatology of EPE length scales, and Figure 5 summarizes these results based on NRM regions. We will describe the seasonal and regional variation in length scale in more detail below; here we note that these seasonal and re-

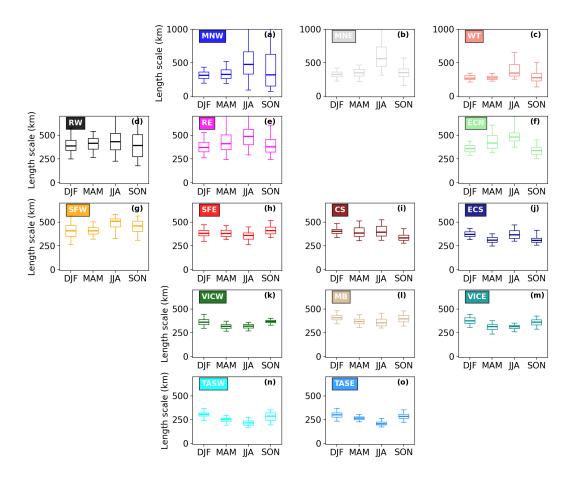


Figure 5. Box plots showing the seasonal length scales of EPEs for stations grouped by NRM regions. Boxes span the 25th–75th percentiles, with the centre line marking the median, and whiskers extending from the 5th to the 95th percentiles. Seasons are shown on the x-axis and estimated length scales on the y-axis, with subplots corresponding to NRM regions. The y-axis is consistent across subplots to enable direct comparison, except for the top row, which shows results for NRM regions in northern Australia. In these regions, higher uncertainty and lower confidence in some estimates arise from low event numbers (< 10; see Figure 6) and sparse station coverage.

gional patterns are coherent, and they do not simply mirror those of P90 thresholds (Figure 2) or mean precipitation (Figure S4).

For reference, we present the number of EPEs used in the semivariogram estimation at each station in each season in Figure 6. The seasonal and regional patterns of the number of EPEs align with those of P90 thresholds (Figure 2), with a higher number of EPEs generally corresponding to higher P90 thresholds. For example, due to the monsoon influence, both Monsoon North West and Monsoon North East exhibit a low number of EPEs during JJA and SON, which largely correspond to the dry season, where the P90 thresholds are also lower. Note that fewer than 10 EPEs were recorded at stations in Monsoon North West during JJA (brown grid boxes in Figure 6). Our sensitivity test (see Figure S5) showed that a low sample size, especially lower than 20, may lead to inaccurate estimates of the climatological length scale shown in Figure 5. We present the results for low EPE numbers for completeness, but any interpretation should account for the potential inaccuracies associated with the small sample size.

We assessed the statistical significance of seasonal variations within each NRM region using the paired Wilcoxon signed-rank test (Wilcoxon, 1945), and seasonal differences between NRM regions using the Mann-Whitney U test (Mann & Whitney, 1947), following approaches similar to Touma et al. (2018) and Tan et al. (2021). These tests were used to evaluate whether median EPE length scales differ significantly across seasons or between regions. The computation of the paired Wilcoxon signed-rank test, Mann-Whitney U test, and p values is conducted using the SciPy package (Virtanen et al., 2020). Results of statistical significance tests for the intraseasonal and intraregional differences are shown in Figure 7 and Figure 8, respectively.

3.1.1 Intraseasonal comparison

Seasonal differences in EPE length scales across Australia's 15 NRM regions reflect the continent's climatic diversity (Figure 7). Most regions exhibit a seasonal cycle with seasonal changes of less than 100 km in spatial extent.

The monsoon-impacted tropical regions (Monsoon North West, Monsoon North East, Wet Tropics) and adjacent East Coast North present a notable paradox: while the number of EPEs is lowest during the dry season (JJA and SON; Figure 6), the median length scales are at their largest (Figure 5a-c). This may be due to a combination of the low number of events and the lower P90 thresholds at tropical stations during JJA, the monsoon inactive phase (Figure 2). Consequently, a storm with a given physical footprint and intensity will exceed the P90 threshold at more stations in the dry season, making its estimated length scale appear larger than an equivalent storm in the wet season. Note that these dry season events are classified as extremes only when seasonal P90 thresholds are applied.

In the eastern regions, Rangeland East and East Coast North exhibit the largest median climatological length scale in JJA (Figure 5e, f), while East Coast South and Central Slope show the largest median climatological length scale in DJF (Figure 5f, i). The southern regions (Southern Flatlands East, Tasmania West, Tasmania East, Murray Basin, Victoria West, and Victoria East) generally show a greater number of EPEs in JJA (Figure 6) and a larger climatological length scale in DJF and SON (Figure 5h-o).

In the western regions (Southern Flatlands West and Rangeland West), both EPE frequency and length scale peak during JJA (Figure 5d, g; Figure 6), coinciding with enhanced frontal precipitation in this season (Prescott et al., 1952; Pook et al., 2012; Raut et al., 2014). In contrast, we find no evidence of similar frontal contributions to EPEs in other southern regions, based on the analysis in Pepler et al. (2020), indicating that the drivers of EPE spatial scales vary across regions.

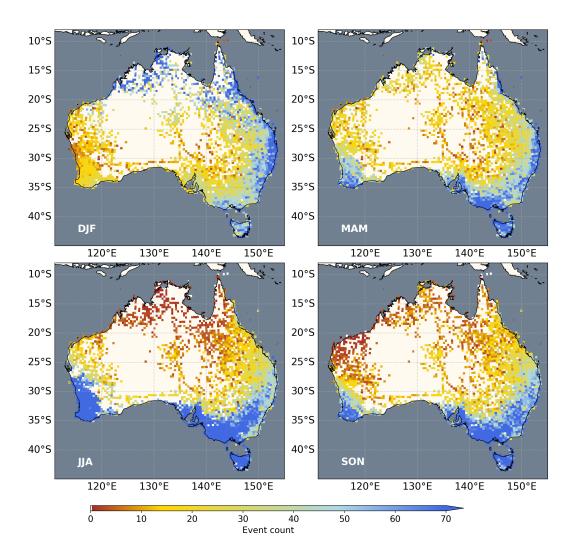


Figure 6. Maps similar to Figure 4, but showing the number of EPEs included at each station in the climatological length scale analysis from 1940 to 2024.

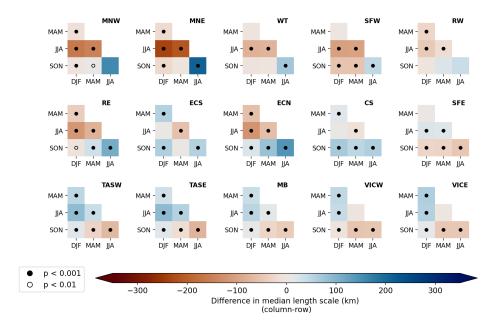


Figure 7. Heatmap showing magnitude and significance of the intraseasonal differences in the median seasonal length scales of the EPEs using the paired Wilcoxon signed-rank test. The shading of each box represents the differences in the median length scale of the column season minus that of the row season. The marker inside each box illustrates the significance level of the paired Wilcoxon signed-rank test as shown in the legend.

3.1.2 Intraregional comparison

Most regional differences in median length scales are highly significant (p < 0.01; Figure 8), confirming that these are robust patterns rather than random chance. Western regions (Southern Flatlands West and Rangeland West) exhibit consistently larger length scales than other regions in all seasons, with the contrast most pronounced in JJA, when variability across the continent is greatest and strongly significant (p < 0.01; Figure 8).

In contrast, Tasmania (TASW and TASE) shows the smallest climatological length scales of any region (Figure 4). A sensitivity test using neighbourhood radii of 300 km and 700 km (not shown) confirmed that these short length scales are not an artifact of the 500-km radius; they remained consistently smaller than those on the Australian mainland across all test cases. Tropical regions (Monsoon North West, Monsoon North East, and Wet Tropics) also display small EPE length scales relative to most other areas (Figure 8). The reduced length scales observed in both Tasmania and the tropical regions are consistent with the findings of the global study by Tan et al. (2021).

3.2 Seasonal length scales of ENSO events

Figures 9, 10, and 11 show the seasonal length scales of EPEs for El Niño and La Niña conditions, along with their differences at each station (La Niña length scale minus El Niño length scale). The displayed results show the median length scales of 1000 bootstrap samples for both ENSO phases. At several stations, the differences in EPE length scales between ENSO phases are statistically significant at the 90% confidence level, indicating that these differences are unlikely to be due to random variability. Since

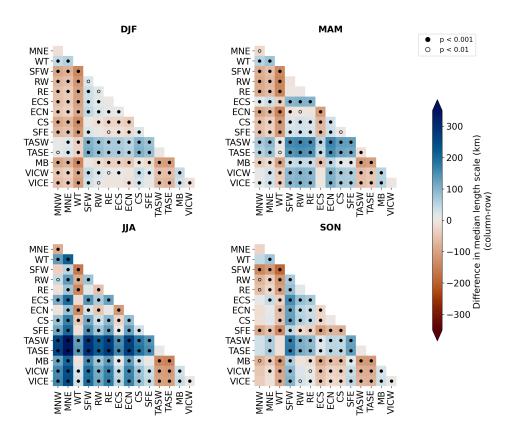


Figure 8. Similar to Figure 7, but for the magnitude and significance of the intraregional differences in the median length scales for each season using the Mann-Whitney U test.

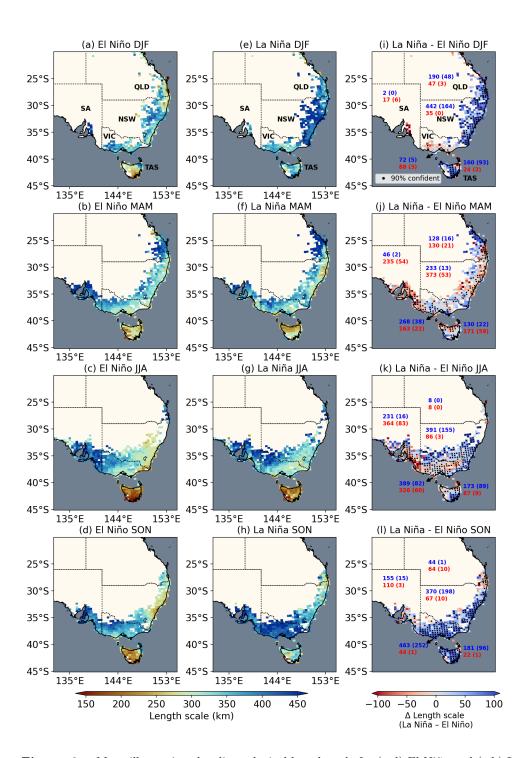


Figure 9. Maps illustrating the climatological length scale for (a-d) El Niño and (e-h) La Niña in eastern and southeastern Australia. The displayed values are median length scales of 1000 bootstrap iterations for each station. Panel (i-l) shows the difference in median length scale between La Niña and El Niño. The locations of Australian states/territories are marked for geographic reference (SA = Southern Australia, QLD = Queensland, NSW = New South Wales, VIC = Victoria, and TAS = Tasmania). For each state or territory within the displayed area, the difference panels (i-l) also indicate the number of stations with larger EPE length scales during La Niña (blue) and El Niño (red). Values in parentheses denote the number of stations with 90% confidence level in the difference in length scales between ENSO phases. Stippling with black dots indicates grid boxes that contain at least one station exceeding the 90% confidence level.

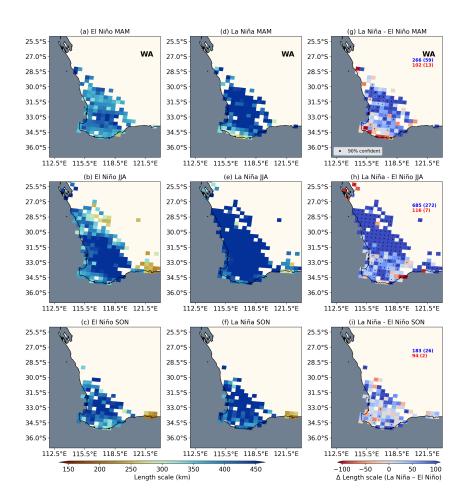


Figure 10. Similar to Figure 9, but for southwestern Australia. Results for DJF are not displayed in this region due to insufficient ENSO events at stations in this region. The displayed region is located in Western Australia (WA).

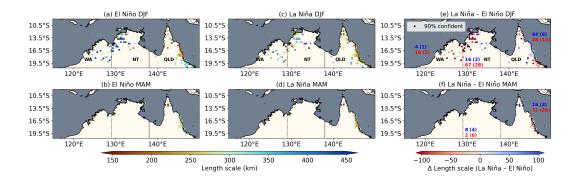


Figure 11. Similar to Figure 9, but for northern Australia. Results for JJA and SON are not shown due to insufficient ENSO events at stations within the displayed region. State/territory boundaries are marked for WA, Northern Territory (NT), and QLD.

several NRM regions have fewer than 10 stations in certain seasons, we present the analysis based on Australian states and territories. Here, we primarily focus on states and territories with sufficient station coverage. For simplicity in geographical reference, the Australian Capital Territory (ACT) is included as part of New South Wales (NSW). Note that Queensland (QLD) in JJA (Figure 9 c, g, k), north Western Australia (WA) in DJF and MAM (Figure 11), and Northern Territory (NT) in MAM (Figure 11 b, d, f) include fewer than 10 stations. Caution should be exercised when interpreting these results.

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While La Niña episodes are associated with wetter conditions than El Niño in eastern Australia with respect to both mean and extreme precipitation (e.g. Nicholls et al., 1997; Risbey et al., 2009; King et al., 2014), the differences in EPE length scales between the two ENSO phases exhibit strong seasonality and vary across regions. In eastern and southeastern Australia, the east coast of Queensland and New South Wales exhibits larger EPE length scales during La Niña in DJF (Figure 9i), with the opposite in MAM (Figure 9i). Moving to JJA and SON, stations in New South Wales still show a larger EPE length scale associated with La Niña (Figure 9k-1). Southeast Queensland shows smaller EPE length scales during La Niña compared to El Niño in SON (Figure 91). The seasonality in Tasmania (TAS) is similar to that in New South Wales, with La Niña linked to a smaller EPE length scale in MAM only. Stations in Victoria (VIC) do not show a clear ENSO-related signal in DJF, but La Niña EPE length scales are generally larger than their El Niño counterparts in other seasons. In Southern Australia (SA), most stations show a larger climatological length scale during El Niño across all seasons, except for SON (Figure 9i-1). In contrast, the opposite pattern is observed in southwestern Australia (Figure 10), where the majority of stations display larger EPE length scales during La Niña compared to El Niño in MAM, JJA, and SON. In northern Australia, only a small number of stations have sufficient samples, but most stations show smaller EPE length scales during La Niña compared to El Niño in DJF (Figure 11e). This pattern is also observed along the northeast coast of Queensland in MAM (Figure 11f).

3.3 Long-term trends in the seasonal length scales

Figure 12 compares the differences of the median values for each NRM region between 1960-1991 and 1992-2023. More details of the seasonal length scales in each NRM region are shown in the boxplots and maps in Figures S6-S7, respectively. Here, we focus primarily on the long-term changes in the medians of climatological length scale in each NRM region. The statistical significance of the differences is evaluated using the Mann-Whitney U test.

Only southwestern Australia (South Flatlands West) and most regions in the east and southeast (East Coast South, East Coast North, Central Slope, South Flatlands East, Tasmania East, Murray Basin, Victoria West, and Victoria East) have a good spatial coverage of stations (Figure S2). Therefore, we mainly focus on these regions for the trends in the climatological length scale. As changes in rain gauge network density may affect length scale estimates, we present results based solely on the 2,070 long-term stations here. We conducted the same analysis that includes all 9861 neighbour stations within neighbourhoods centred on the long-term stations. The results from both the long-term and 9861 neighbour station sets are largely consistent (not shown).

The trend analysis of EPE length scales reveals geographically coherent patterns across Australia, with most of the observed changes being statistically significant (Figure 12). A notable increasing trend is seen across the eastern and southeastern regions (Victoria West, Victoria East, East Coast South, and East Coast North) during SON, while a decreasing trend is common in MAM. In contrast, the Tasmanian regions (TASW and TASE) are the main areas where the changes are not statistically significant, while they share a similar trend of decreasing in DJF and JJA, and increasing in SON. Fur-

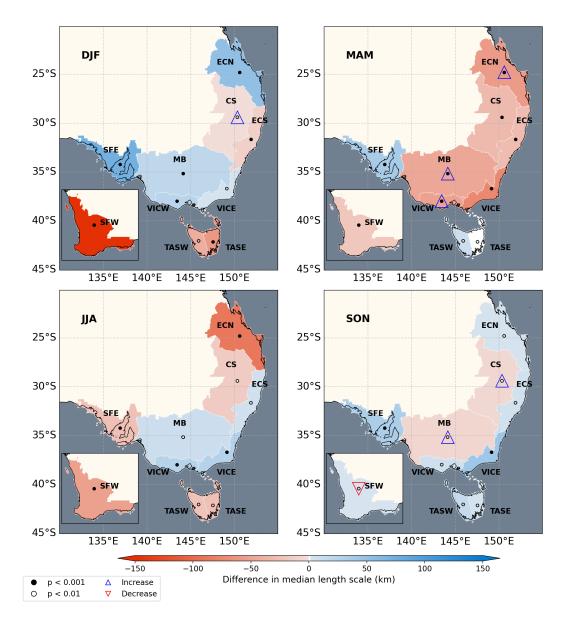


Figure 12. Maps showing magnitude and significance of the differences in the median seasonal length scales of the EPEs for each NRM region using the Mann-Whitney U test for DJF (top left), MAM (top right), JJA (bottom left), and SON (bottom right). The shading of each region represents the differences of the median values in 1992-2023 minus those of 1960-1991. Circle markers at region centroids illustrate the p values of the Mann-Whitney U test. The triangle marker shows where excluding the Millennium Drought reverses the trend. Each panel includes an inset map of SFW (bottom left). Results for MNW, MNE, WT, RW, and RE are not shown due to insufficient stations.

ther west, Southern Flatlands East shows an increase in all seasons except JJA, while Southern Flatlands West exhibits a general decreasing trend except SON.

Southern and southeastern Australia (all displayed regions in Figure 12 except for Southern Flatlands West and East Coast North) experienced a prolonged dry period before 2010, commonly known as the "Millennium Drought" (Van Dijk et al., 2013). To isolate the influence of this drought on changes in the EPE length scales, we conducted an additional analysis using two periods that exclude the drought years: 1960–1973 and 2010–2023. For most regions, this analysis yields consistent results (see Figure 12). Therefore, we only highlight the specific cases where excluding the Millennium Drought produced a significant change.

The influence of the Millennium Drought is evident in some seasons in Southern Flatlands West, Murray Basin, Central Slopes, East Coast South, East Coast North, Victoria West, and Tasmanian East. In the far west (Southern Flatlands West), the trend in SON reverses from increasing to decreasing when the drought is excluded. Moving to the east, the trend in Victoria West and Murray Basin in MAM flips to increasing, while the increasing trend is weak (< 18 km; not shown). For East Coast North, such a reversal to an increasing trend is observed during MAM. In Murray Basin and Central Slope, the trends in SON flip from a decreasing to an increasing trend.

The above patterns in changes to the length scale of EPEs do not typically correspond to those in mean precipitation (Figure S8) or the P90 thresholds (Figure S9), with two notable exceptions. First, the strong reduction in EPE length scales in Southern Flatlands West aligns with a documented drying trend in southwestern Australia (Dey et al., 2019). This comparison between extreme events and mean trends is relevant given the established strong relationship between them (Nishant & Sherwood, 2021). Second, a decreasing trend is shared by both median length scales and mean precipitation during MAM across southeastern and eastern Australia (Victoria West and East, Tasmania West and East, Murray Basin, Central Slope, and East Coast South and North; Figure S8). The implications of these results are explored in the next section.

4 Discussion

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A more humid large-scale environment is often associated with greater precipitation amounts and broader rainfall coverage (e.g. environments with higher mid-level relative humidity and total precipitable water; Zhou et al., 2013), suggesting that both the intensity and spatial scale of extremes may increase under such conditions. However, our results suggest a more complex picture with the spatial scales of EPEs. Unlike mean precipitation, which generally shows a positive correlation with La Niña conditions across Australia (Risbey et al., 2009), the response of EPE length scales does not follow the same pattern. In some instances, the patterns align with expectations. For example, the larger length scales in southeastern Australia during Spring (SON) under La Niña are consistent with conditions favourable for rain-producing cyclones (Gillett et al., 2023), which produce large-scale EPEs. However, in other cases, the relationship is inverted, such as the larger length scales observed during El Niño in South Australia in JJA (Figure 9k). To ensure the robustness of our analysis, we verified that the selected ENSO events are not dominated by any single year (i.e., a particularly wet El Niño or a particularly dry La Niña) and that, consistent with established patterns, most stations coincide with higher seasonal total and mean precipitation in each La Niña year. Understanding the mechanisms that drive contrasting responses of EPE intensity and EPE length scale to ENSO variations is an important avenue for future research.

The lack of correspondence between length scale patterns and mean precipitation is echoed in the trend analysis. Several eastern and southeastern regions (Victoria and East Coast) show a reduction in MAM and an increase in SON. Since mean precipita-

tion shows no significant change from MAM to SON (Figure S8), the difference is likely driven by the spatial characteristics of the underlying weather systems rather than total rainfall.

We propose two possible explanations for the above finding. The increasing length scales in SON could be due to a seasonal shift in large-scale circulation, causing the synoptic systems that produce widespread precipitation to occur earlier in the year and/or extend further inland. While there are weak trends in mean precipitation in coastal regions, there are increasing trends in mean precipitation (Figure S8) and the P90 threshold (Figure S9) observed at inland stations that may be indicative of a broader inland shift in precipitation patterns. Confirming whether these contrasting shifts in EPE length scale and EPE intensity can be explained by shifts in large-scale circulation will require further investigation into the relationship between large-scale atmospheric conditions and EPE spatial characteristics.

A second explanation focuses on the decreasing length scales in MAM. While a warming climate generally favours more convective systems that concentrate rainfall over smaller areas (e.g., Peleg et al., 2018; Lochbihler et al., 2019), the fact that this trend is most evident in MAM suggests a role for seasonally specific storm-type and large-scale circulation changes. Further work is required to determine if such a shift in rain type is occurring in the relevant regions, especially using higher spatial resolution observations that can resolve storm morphology, such as precipitation estimates from radar or satellites.

While the semivariogram approach is useful to estimate the climatological length scales of EPEs, it assumes isotropy, meaning that the spatial correlation of precipitation decays uniformly in all directions. This isotropy assumption has two key limitations. First, the method does not account for topographic influences. For example, when a large weather system interacts with a mountain range, it can produce intense orographic precipitation on the windward side and a pronounced rain shadow on the leeward side. Our analysis would interpret this as a smaller-scale EPE, failing to distinguish it from a genuinely localised convective storm, even though their physical drivers differ fundamentally. Second, an isotropic framework cannot capture anisotropic (e.g., elongated) precipitation footprints, such as a long but narrow swath of heavy rain produced by a fast-moving, small-scale storm. The isotropic method averages this footprint into a single length scale, losing important information about the storm's shape and motion. Accounting for anisotropy is important because the footprint's geometry reflects an EPE's physical drivers and affects its hydrological consequences, such as whether it causes localised or widespread flooding.

A further limitation is in the choice of the theoretical semivariogram model, which can influence the estimated length scales. In this study, the selection was made heuristically using synthetic data, while the optimal model for a more accurate estimation of length scales may vary across different parts of Australia. Future work should incorporate improved model selection procedures alongside methods that capture anisotropic precipitation patterns to better represent the spatial characteristics of extreme precipitation events (see, e.g., Johnston et al., 2001; Niemi et al., 2014; Saunders et al., 2021; Verbovšek, 2024, and references therein).

5 Conclusions

We analysed the long-term climatological length scale of EPEs in Australia using the relaxed neighbourhood semivariogram approach. BoM station data from 1940 to 2024 were used to estimate the seasonal length scale of EPEs. Australia was divided into 15 regions based on the NRM clusters. The seasonal analysis reveals diverse characteristics in EPE length scales across Australia, with geographically coherent patterns emerging among adjacent regions. The seasonality differs across climate zones. Monsoon-impacted

regions exhibit a visible separation between dry and wet seasons in the number of EPEs, while other regions display a clear seasonal cycle in climatological length scales. Although Tasmania is associated with the smallest length scale, it still shares geographical similarity in the seasonal cycle of length scales with nearby regions. The southern regions coincide with the smallest length scale in JJA, while the eastern regions present the largest length scale in JJA. The paired Wilcoxon signed-rank test and Mann-Whitney U test were carried out to show that the variations in the median climatological length scale are statistically significant across seasons and regions.

Based on the SOI, we present the first climatological analysis of EPE spatial scales in Australia during different ENSO phases. Although La Niña conditions are typically associated with increased precipitation in eastern Australia, our results show that EPEs are not uniformly larger in spatial scale during La Niña compared to El Niño across all seasons and all regions. In MAM, the La Niña EPEs mostly display a smaller length scale than their El Niño counterparts on the east coast, while this pattern is reversed in other seasons.

With the availability of long-term data, we also examined changes in the climato-logical length scale by splitting the analysis into two periods, one from 1960 to 1991 and the other from 1992 to 2023. Regardless of the Millennium drought, the analysis reveals a reduction in the spatial scale of EPEs in several regions in southeastern and eastern Australia in MAM, but an increase in SON. Crucially, these trends, along with the ENSO relationships, differ from those observed in mean precipitation and intensity of extreme precipitation. This suggests that the mechanisms governing the spatial scale of extremes are distinct from those controlling mean precipitation, and that changes in mean precipitation do not necessarily imply corresponding changes in the spatial extent of extreme events.

In summary, this study provides a valuable climatological reference for the spatial extent of EPEs in Australia. By establishing how EPE spatial scales vary, we lay the foundation for future work to investigate the underlying physical mechanisms. For example, while our results suggest a link between frontal systems and large-scale EPEs in southwestern Australia in JJA, whether such systems are directly responsible for large-scale extremes in this region remains an open question. Similarly, we reveal a complex relationship between EPE spatial scales and climate drivers like ENSO. This complexity highlights that improving projections of extreme events in a changing climate requires more than just understanding changes in mean rainfall. An important future direction is therefore to go beyond the climatological analysis here and investigate how day-to-day large-scale atmospheric conditions modulate the spatial characteristics of EPEs, and the physical mechanisms driving this modulation. Such work will strengthen the physical basis of our projections for how extreme rainfall will change in a warming climate.

Data and software availability

The BoM daily rain gauge data are freely available from BoM's weather station directory (http://www.bom.gov.au/climate/data/stations/; last accessed: 26 August 2025). The SOI was obtained from the BoM (http://www.bom.gov.au/climate/enso/soi/; last accessed: 26 August 2025). The code for semivariogram estimation and the gridded maps used to create the figures in this paper are archived by Lin (2025). Figures in this paper were produced using the scientific colour maps of Crameri et al. (2020).

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References

- Ashouri, H., Hsu, K.-L., Sorooshian, S., Braithwaite, D. K., Knapp, K. R., Cecil, L. D., . . . Prat, O. P. (2015). Persiann-cdr: Daily precipitation climate data record from multisatellite observations for hydrological and climate studies. Bulletin of the American Meteorological Society, 96(1), 69–83.
 - Berghuijs, W. R., Allen, S. T., Harrigan, S., & Kirchner, J. W. (2019). Growing spatial scales of synchronous river flooding in europe. Geophysical Research Letters, 46(3), 1423–1428.
 - Bevacqua, E., Shepherd, T. G., Watson, P. A., Sparrow, S., Wallom, D., & Mitchell, D. (2021). Larger spatial footprint of wintertime total precipitation extremes in a warmer climate. Geophysical Research Letters, 48(8), e2020GL091990.
 - Bevacqua, E., Zappa, G., & Shepherd, T. G. (2020). Shorter cyclone clusters modulate changes in european wintertime precipitation extremes. *Environmental Research Letters*, 15(12), 124005.
 - Chang, W., Stein, M. L., Wang, J., Kotamarthi, V. R., & Moyer, E. J. (2016). Changes in spatiotemporal precipitation patterns in changing climate conditions. *Journal of Climate*, 29(23), 8355–8376.
 - Crameri, F., Shephard, G. E., & Heron, P. J. (2020). The misuse of colour in science communication. *Nature communications*, 11(1), 5444.
 - Cressie, N. (2015). Statistics for spatial data. John Wiley & Sons.
 - CSIRO. (2015). Climate change in australia, technical report: Projections for australia's nrm regions. Commonwealth Scientific and Industrial Research Organisation and Australian Bureau of Meteorology Marine and Atmospheric Research Melbourne, Victoria.
 - Dey, R., Lewis, S. C., Arblaster, J. M., & Abram, N. J. (2019). A review of past and projected changes in australia's rainfall. Wiley Interdisciplinary Reviews: Climate Change, 10(3), e577.
 - Drosdowsky, W. (1993). An analysis of australian seasonal rainfall anomalies: 1950–1987. i: Spatial patterns. *International Journal of Climatology*, 13(1), 1–30.
 - Dwyer, J., & O'Gorman, P. (2017). Changing duration and spatial extent of midlatitude precipitation extremes across different climates. *Geophysical Research Letters*, 44(11), 5863–5871.
 - Gillett, Z., Taschetto, A., Holgate, C., & Santoso, A. (2023). Linking enso to synoptic weather systems in eastern australia. *Geophysical Research Letters*, 50(15), e2023GL104814.
 - Guinard, K., Mailhot, A., & Caya, D. (2015). Projected changes in characteristics of precipitation spatial structures over north america. *International Journal of Climatology*, 35(4).
 - He, J., Li, S., Wang, B., Zhang, L., & Duan, K. (2025). Quantifying the impacts of enso on australian summer rainfall extremes during 1960–2020. *Journal of Hydrology*, 132834.
 - Hoegh-Guldberg, O., Jacob, D., Taylor, M., Bindi, M., Brown, S., Camilloni, I. A., ... Zhou, G. (2018). Impacts of 1.5 °c global warming on natural and human systems. An IPCC Special Report on the impacts of global warming of 1.5 °C above pre-industrial levels and related global greenhouse gas emission pathways, in the context of strengthening the global response to the threat of climate change, sustainable development, and efforts to eradicate poverty, 175–311.

Huang, A. T., Gillett, Z. E., & Taschetto, A. S. (2024). Australian rainfall increases during multi-year la niña. Geophysical Research Letters, 51(9), e2023GL106939.

- Johnston, K., Ver Hoef, J. M., Krivoruchko, K., & Lucas, N. (2001). *Using arcgis geostatistical analyst* (Vol. 380). Esri Redlands.
- Jongman, B., Hochrainer-Stigler, S., Feyen, L., Aerts, J. C., Mechler, R., Botzen, W. W., ... Ward, P. J. (2014). Increasing stress on disaster-risk finance due to large floods. *Nature Climate Change*, 4(4), 264–268.
- Khouakhi, A., Villarini, G., & Vecchi, G. A. (2017). Contribution of tropical cyclones to rainfall at the global scale. *Journal of Climate*, 30(1), 359–372.
- King, A. D., Klingaman, N. P., Alexander, L. V., Donat, M. G., Jourdain, N. C., & Maher, P. (2014). Extreme rainfall variability in australia: Patterns, drivers, and predictability. *Journal of Climate*, 27(15), 6035–6050.
- King, A. D., Pitman, A. J., Henley, B. J., Ukkola, A. M., & Brown, J. R. (2020). The role of climate variability in australian drought. *Nature Climate Change*, 10(3), 177–179.
- Lin, D. (2025, August). Semi-variogram_AU: Semivariogram v1.1 with gridded map. Zenodo. doi: 10.5281/zenodo.16946013
- Lochbihler, K., Lenderink, G., & Siebesma, A. P. (2017). The spatial extent of rainfall events and its relation to precipitation scaling. Geophysical Research Letters, 44(16), 8629–8636.
- Lochbihler, K., Lenderink, G., & Siebesma, A. P. (2019). Response of extreme precipitating cell structures to atmospheric warming. *Journal of Geophysical Research: Atmospheres*, 124(13), 6904–6918.
- Mann, H. B., & Whitney, D. R. (1947). On a test of whether one of two random variables is stochastically larger than the other. The annals of mathematical statistics, 50–60.
- McBride, J. L., & Nicholls, N. (1983). Seasonal relationships between australian rainfall and the southern oscillation. *Monthly Weather Review*, 111(10), 1998–2004.
 - McGregor, S., Gallant, A., & van Rensch, P. (2024). Quantifying ensos impact on australia's regional monthly rainfall risk. Geophysical Research Letters, 51(6), e2023GL106298.
 - Montoya Duque, E., Huang, Y., May, P., & Siems, S. (2023). An evaluation of imerg and era5 quantitative precipitation estimates over the southern ocean using shipborne observations. *Journal of Applied Meteorology and Climatology*, 62(11), 1479–1495.
 - Müller, S., Schüler, L., Zech, A., & Heße, F. (2022). GSTools v1.3: a toolbox for geostatistical modelling in python. Geoscientific Model Development, 15(7), 3161-3182. Retrieved from https://gmd.copernicus.org/articles/15/3161/2022/ doi: 10.5194/gmd-15-3161-2022
 - Nicholls, N., Drosdowsky, W., & Lavery, B. (1997). Australian rainfall variability and change. Weather, 52(3), 66–72.
 - Niemi, T. J., Kokkonen, T., & Seed, A. W. (2014). A simple and effective method for quantifying spatial anisotropy of time series of precipitation fields. *Water Resources Research*, 50(7), 5906–5925.
 - Nishant, N., & Sherwood, S. C. (2021). How strongly are mean and extreme precipitation coupled? *Geophysical Research Letters*, 48(10), e2020GL092075.
 - O'Gorman, P. A., & Schneider, T. (2009). The physical basis for increases in precipitation extremes in simulations of 21st-century climate change. *Proceedings* of the National Academy of Sciences, 106(35), 14773–14777.
 - O'Gorman, P. A. (2015). Precipitation extremes under climate change. Current climate change reports, 1, 49–59.
- Paciorek, C. J., Stone, D. A., & Wehner, M. F. (2018). Quantifying statistical uncertainty in the attribution of human influence on severe weather. Weather and

climate extremes, 20, 69–80.

- Peleg, N., Marra, F., Fatichi, S., Molnar, P., Morin, E., Sharma, A., & Burlando, P. (2018). Intensification of convective rain cells at warmer temperatures observed from high-resolution weather radar data. *Journal of Hydrometeorology*, 19(4), 715–726.
- Pepler, A. S., Dowdy, A. J., Van Rensch, P., Rudeva, I., Catto, J. L., & Hope, P. (2020). The contributions of fronts, lows and thunderstorms to southern australian rainfall. *Climate Dynamics*, 55, 1489–1505.
- Ponukumati, P., Mohammed, A., & Regonda, S. (2023). Insights on satellite-based imerg precipitation estimates at multiple space and time scales for a developing urban region in india. *Journal of Hydrometeorology*, 24(6), 977–996.
- Pook, M. J., Risbey, J. S., & McIntosh, P. C. (2012). The synoptic climatology of cool-season rainfall in the central wheatbelt of western australia. *Monthly Weather Review*, 140(1), 28–43.
- Prescott, J. A., Collins, J. A., & Shirpurkar, G. (1952). The comparative climatology of australia and argentina. *Geographical Review*, 42(1), 118–133. doi: 10.2307/211256
- Raut, B. A., Jakob, C., & Reeder, M. J. (2014). Rainfall changes over southwestern australia and their relationship to the southern annular mode and enso. *Journal of Climate*, 27(15), 5801–5814.
- Risbey, J. S., Pook, M. J., McIntosh, P. C., Wheeler, M. C., & Hendon, H. H. (2009). On the remote drivers of rainfall variability in australia. *Monthly Weather Review*, 137(10), 3233–3253.
- Saunders, K., Stephenson, A., & Karoly, D. (2021). A regionalisation approach for rainfall based on extremal dependence. *Extremes*, 24, 215–240.
- Schaller, N., Kay, A. L., Lamb, R., Massey, N. R., Van Oldenborgh, G. J., Otto, F. E., . . . others (2016). Human influence on climate in the 2014 southern england winter floods and their impacts. *Nature Climate Change*, 6(6), 627–634.
- Schumacher, R. S. (2017). Heavy rainfall and flash flooding. In Oxford research encyclopedia of natural hazard science.
- Simpson, J., Kummerow, C., Tao, W. K., & Adler, R. F. (1996). On the tropical rainfall measuring mission (trmm). *Meteorology and Atmospheric physics*, 60, 19–36.
- Tan, X., Wu, X., & Liu, B. (2021). Global changes in the spatial extents of precipitation extremes. *Environmental Research Letters*, 16(5), 054017.
- Tansey, E., Marchand, R., Protat, A., Alexander, S. P., & Ding, S. (2022). Southern ocean precipitation characteristics observed from cloudsat and ground instrumentation during the macquarie island cloud & radiation experiment (micre): April 2016 to march 2017. *Journal of Geophysical Research: Atmospheres*, 127(5), e2021JD035370.
- Touma, D., Michalak, A. M., Swain, D. L., & Diffenbaugh, N. S. (2018). Characterizing the spatial scales of extreme daily precipitation in the united states. Journal of Climate, 31(19), 8023–8037.
- Touma, D., Stevenson, S., Camargo, S. J., Horton, D. E., & Diffenbaugh, N. S. (2019). Variations in the intensity and spatial extent of tropical cyclone precipitation. *Geophysical Research Letters*, 46(23), 13992–14002.
- Tozer, C. R., Risbey, J. S., Monselesan, D. P., Pook, M. J., Irving, D., Ramesh, N., ... Squire, D. T. (2023). Impacts of enso on australian rainfall: What not to expect. *Journal of Southern Hemisphere Earth Systems Science*, 73(1), 77–81.
- Trenberth, K. E. (1984). Signal versus noise in the southern oscillation. *Monthly Weather Review*, 112(2), 326–332.
- Van Dijk, A. I., Beck, H. E., Crosbie, R. S., De Jeu, R. A., Liu, Y. Y., Podger,
 G. M., ... Viney, N. R. (2013). The millennium drought in southeast australia
 (2001–2009): Natural and human causes and implications for water resources,
 ecosystems, economy, and society. Water Resources Research, 49(2), 1040–

1057.

- Verbovšek, T. (2024). Analysis of karst surface anisotropy using directional semivariograms, slovenia. *Pure and Applied Geophysics*, 181(1), 233–245.
- Virtanen, P., Gommers, R., Oliphant, T. E., Haberland, M., Reddy, T., Cournapeau, D., ... SciPy 1.0 Contributors (2020). SciPy 1.0: Fundamental Algorithms for Scientific Computing in Python. Nature Methods, 17, 261–272. doi: 10.1038/s41592-019-0686-2
- Warren, R. A., Jakob, C., Hitchcock, S. M., & White, B. A. (2021). Heavy versus extreme rainfall events in southeast australia. *Quarterly Journal of the Royal Meteorological Society*, 147(739), 3201–3226.
- Wasko, C., Sharma, A., & Westra, S. (2016). Reduced spatial extent of extreme storms at higher temperatures. Geophysical Research Letters, 43(8), 4026–4032
- Webster, R., & Oliver, M. A. (2007). Geostatistics for environmental scientists. John Wiley & Sons.
- White, B., Jakob, C., & Reeder, M. (2022). Fundamental ingredients of australian rainfall extremes. *Journal of Geophysical Research: Atmospheres*, 127(17), e2021JD036076.
- Wilcoxon, F. (1945). Individual comparisons by ranking methods. *Biometrics Bulletin*, 1(6), 80–83. doi: 10.2307/3001968
- Zhou, Y., Lau, W. K., & Liu, C. (2013). Rain characteristics and large-scale environments of precipitation objects with extreme rain volumes from trmm observations. *Journal of Geophysical Research: Atmospheres*, 118(17), 9673–9689.

Supporting Information for "Mapping the Spatial Scales of Australian Extreme Precipitation Using Daily Rain Gauges"

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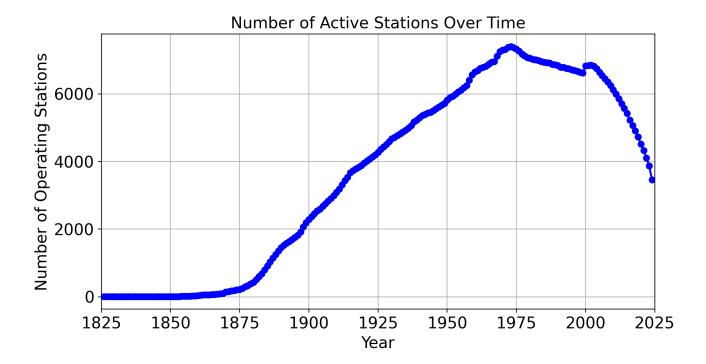


Figure S1. Number of active BoM stations in Australia over time.

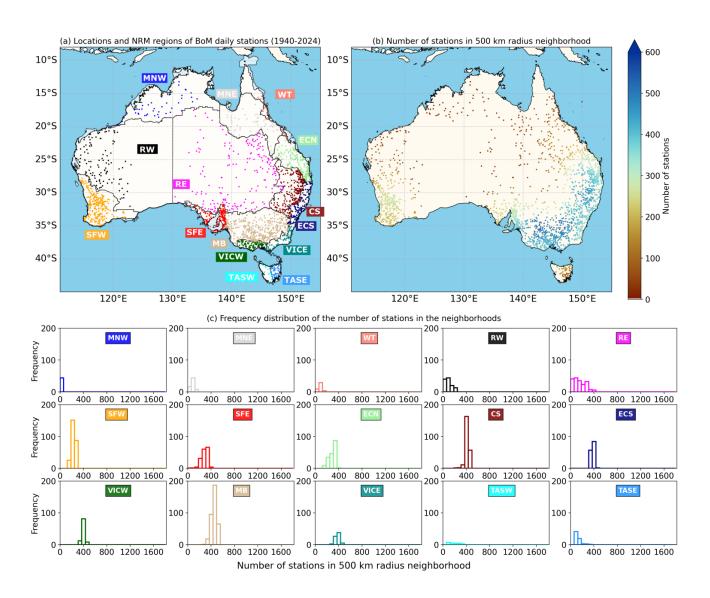


Figure S2. Similar to Figure 1, but for the long-term stations only.

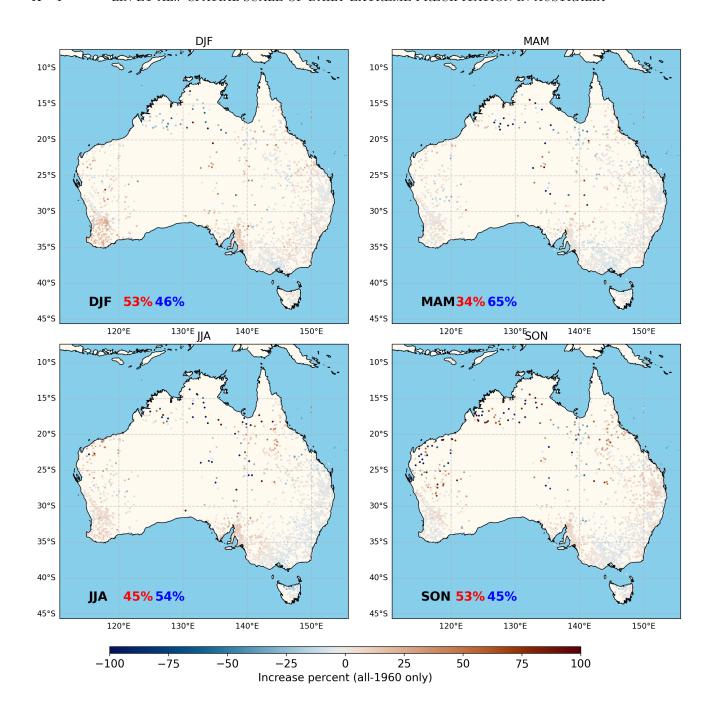


Figure S3. The percentage difference in the climatological length scales estimated using all neighbour stations operated for more than 20 years but not throughout 1960-2023, minus those estimated only using the long-term stations in the 500 km neighbourhood. The red numbers in each panel indicate the percentage of stations with a larger length scale when all neighbour stations are used, and the blue numbers in each panel are the percentage of stations with a smaller length scale when only the long-term stations are used as neighbour stations.

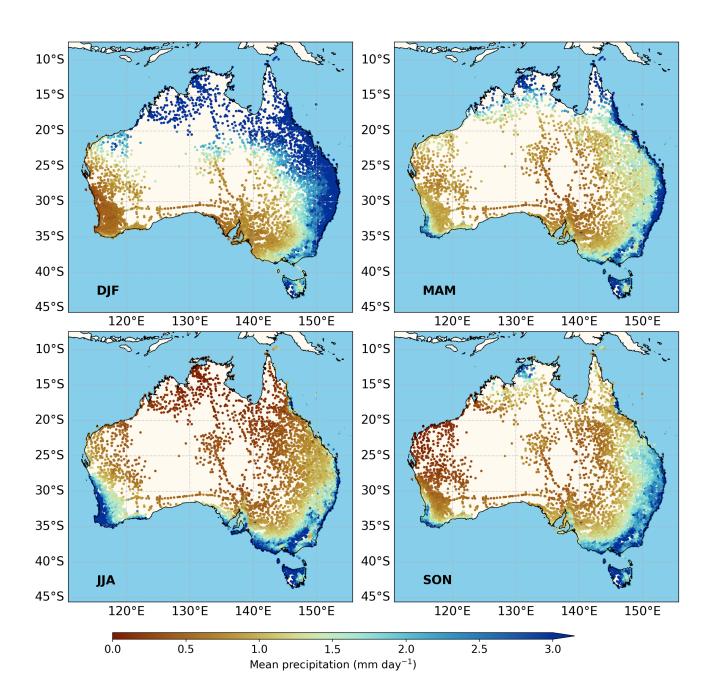


Figure S4. Similar to Figure 2, but for the mean precipitation of all measurements for BoM daily stations with more than 20 years of continuous records between 1940 and 2024.

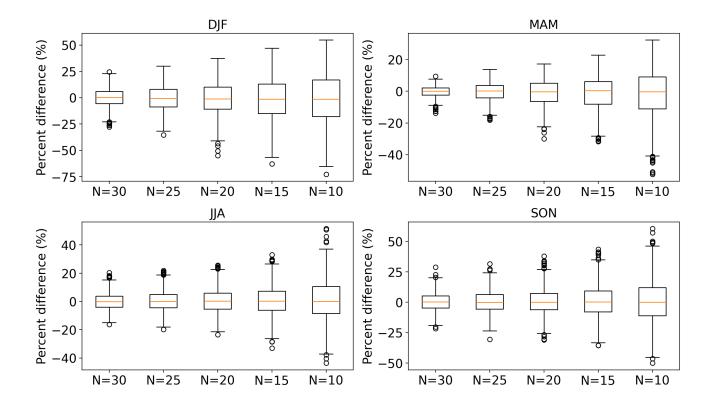


Figure S5. Sensitivity test of the number of EPEs included in the seasonal length scale estimate. N is the number of EPEs. The test was done by randomly selecting 30, 25, 20, 15, and 10 EPEs for each season. The test was repeated 1000 times. The percentage difference is the subsample results minus the reference length scale, which used all EPE days for each season. The percentage difference starts to exceed 50% when $N \le 20$. The results shown here are for the same station shown in Figure 3.

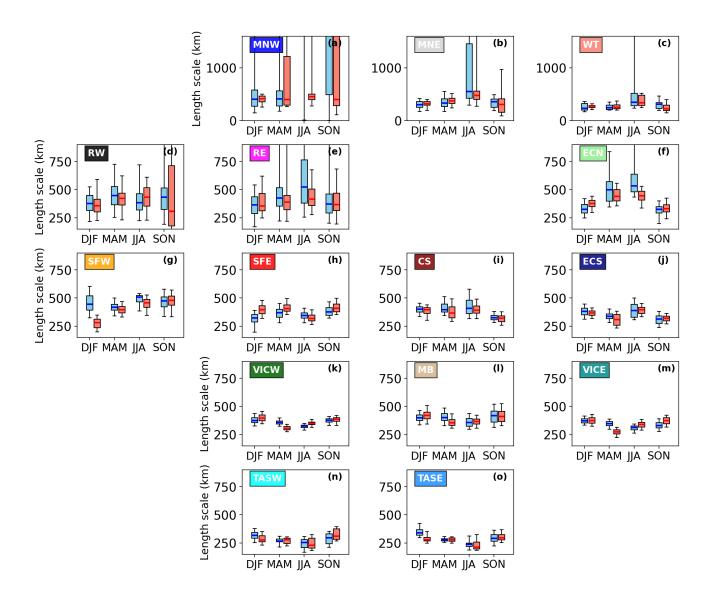


Figure S6. Box plots comparing the spread of the climatological length scales for each NRM region for the two periods of 1960-1991 (blue boxes in each panel) and 1992-2023 (red boxes in each panel). Note the difference in the y-axis scale of the panels in the top row. Note that a low number (< 20) of EPEs were recorded for stations in MNW, MNE, and RW in JJA and/or SON.

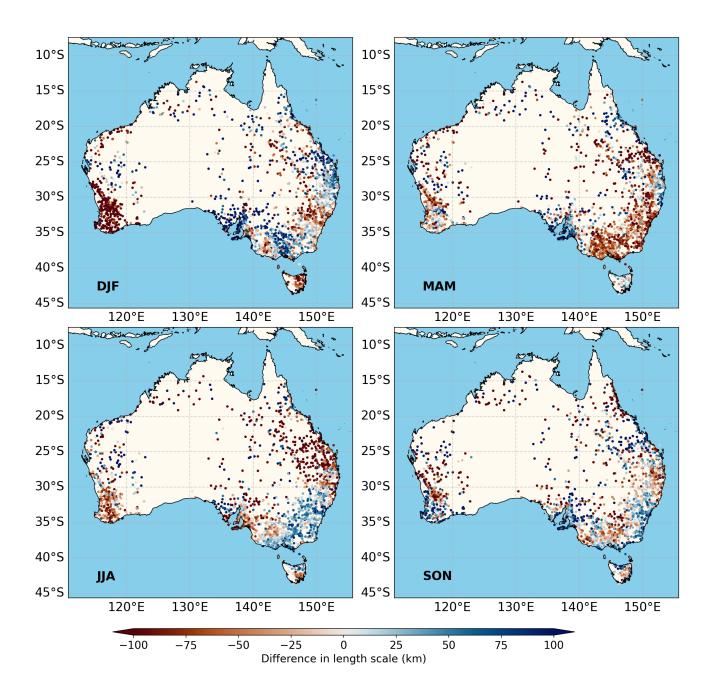


Figure S7. Similar to Figure S4, but for the differences in the climatological length scales between 1960-1991 and 1992-2023.

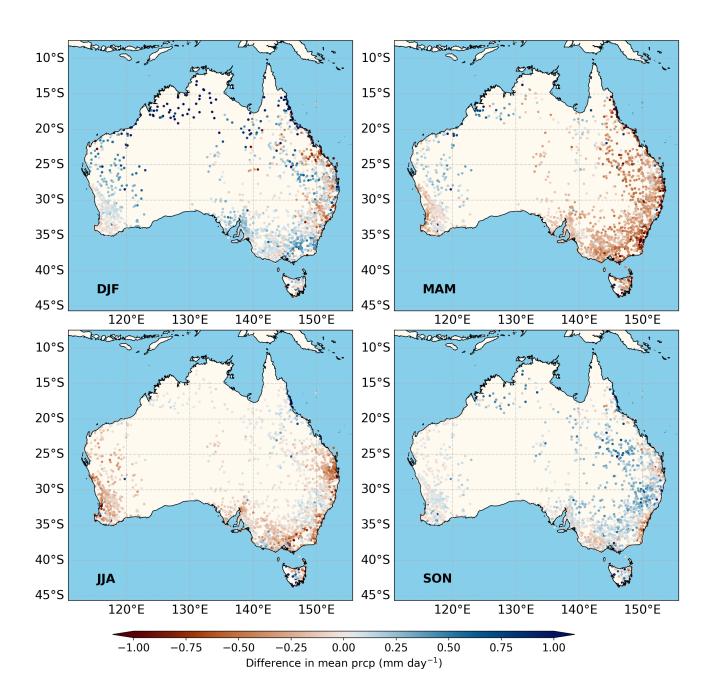


Figure S8. Similar to Figure S7, but for the differences in the daily mean precipitation between 1960-1991 and 1992-2023.

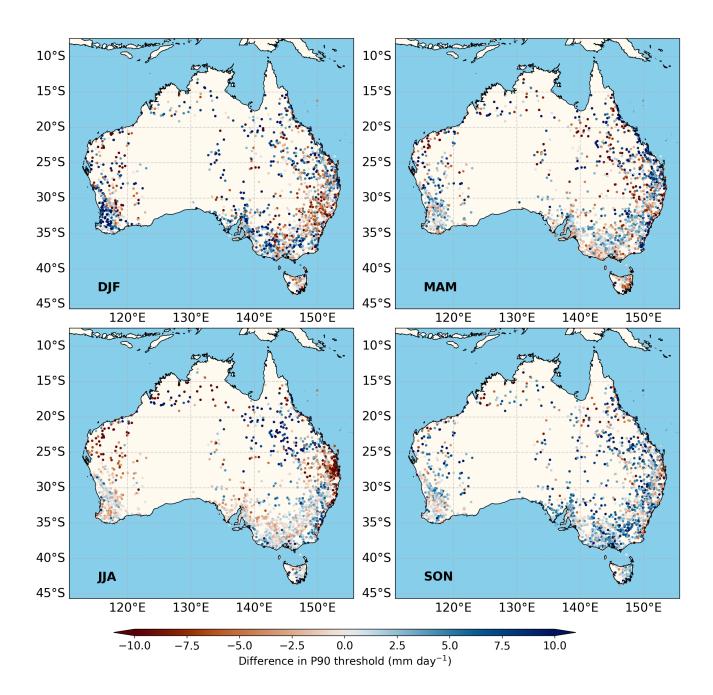


Figure S9. Similar to Figure S7, but for the differences in the P90 thresholds between 1960-1991 and 1992-2023.